<u>Exam</u>



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Info

- In submitting the solutions there is no need to rephrase the problem. "Solution for 1a" is sufficient.
- The submission format for explanations and plots is a PDF file. Also, include any and all software scripts used to establish your answer(s) and/or produce plots.
- Working in groups or any communication about the problems is prohibited. Using the internet as a resource is encouraged, but soliciting any help is also prohibited.
- Please do not zip the files for submission
- Some questions have multiple parts. For full credit, all parts must be done.

- There is a file posted online which has 5 columns, each representing a physical observable of interest generated from some underlying function. There are thousands of entries, i.e. rows.
 - http://www.nbi.dk/~koskinen/Teaching/
 AdvancedMethodsInAppliedStatistics2016/data/Exam_Prob1.txt
 - The first 3 variables/columns are independent distributions with no correlation to the other variables
 - The last two variables/columns are unused
 - Be mindful about accounting for truncated ranges as well as likelihood functions that have periodic components which will create local minima/maxima

Lists of Distributions

- The data in each column is produced from one of the functions shown at right
- Note that these functions may be unnormalized
 - Hint: Some will require a normalization to convert them to probability distribution functions

$$\frac{1}{x+5}\sin(ax)$$

$$\sin(ax) + 1$$

$$\sin(ax^2)$$

$$\sin(ax+1)^2$$

$$x\tan(x)$$

$$1 + ax + bx^2$$

$$a + bx$$

$$e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

$$f(k) \propto \frac{\binom{n}{k} p^k (1-p)^{n-k} \quad binomial}{\binom{k}{k}} \times \frac{\lambda^k e^{-\lambda}}{k!} \qquad poisson$$

$$\frac{-1}{\ln(1-p)} \frac{p^k}{k} \quad logarithmic$$

Problem 1a

- Using the separate data from the first three columns, identify
 the function on slide 4 from which each was generated and
 find the best-fit values for that distribution using the <u>maximum</u>
 <u>likelihood method</u>
 - E.g. if $f(x)=\sin(ax+b)*\exp(-x+c)+x/k!$ were one of the functions, then find the best-fit values for a, b, c, and k
 - Degeneracies exist, e.g. sin(x)=cos(a+x), which can produce functionally identical data distributions
 - Any function, with associated best-fit parameters which is <u>statistically</u> compatible with the data in the files will be accepted as a proper solution.
 Only one is necessary.
- Data in columns 1 and 2 have artificially truncated ranges
 - Column 1 is only sampled in the independent variable from 0 to 2
 - Column 2 is only sampled from -1 to 1
 - Column 3 is not truncated

Problem 1b

- Plot the data and the corresponding best-fit function on the same plots
 - 3 separate 1-dimension plots
 - Plot as a function of the independent variable
 - Histogram the data, and scale the best-fit function to be 'reasonable' so that the features of both the data and best-fit function can be visually compared

• A cancer study in 1991 conducted in Wisconsin collected data from ~700 patients. There were 9 variables associated with the digitized image of a fine needle aspirate biopsy sample of a tissue mass. Each variable has discrete values from 1-10. There was also the patient identifier (code number) and whether the sample mass was ultimately benign or malignant.

# Attribute	Domain/Range
1. Sample code numbe	er id number
2. Clump Thickness	1 - 10
3. Uniformity of Cell Si	ze 1 - 10
4. Uniformity of Cell Sh	nape 1 - 10
5. Marginal Adhesion	1 - 10
6. Single Epithelial Cell Size 1 - 10	
7. Bare Nuclei	1 - 10
8. Bland Chromatin	1 - 10
9. Normal Nucleoli	1 - 10
10. Mitoses	1 - 10
11. Class:	(2 for benign, 4 for malignant)

Problem 2a

- There are two files online: training data and blind data
 - The following training data includes the aforementioned variables as well as whether the biopsy was benign or malignant (www.nbi.dk/ ~koskinen/Teaching/AdvancedMethodsInAppliedStatistics2016/data/breast-cancer-wisconsin_train-test.txt)
 - The following data includes the same variables, but the information of whether the biopsy was benign or malignant has been removed, i.e. a blind sample, (www.nbi.dk/~koskinen/Teaching/
 AdvancedMethodsInAppliedStatistics2016/data/breast-cancer-wisconsin_mod_real.txt)
- Using some method (straight cuts, support vector machine, boosted decision tree, etc.) and the training data, come up with a classification algorithm which uses the 9 variables to identify malignant and benign tissue samples
 - Note: the separate variables can be assumed to have the same features and shape between the training and blind sample

Problem 2a cont.

- With a developed algorithm, run the classifier over the training sample and calculate the efficiency of identifying a malignant mass
 - It is possible to get 100% efficiency, provided the method is overtrained
 - But, you will have to use the same settings for classifying the blind sample in Problem 2b
- Calculate the overall classification efficiency for the whole training sample
 - (classified_true_malignant + classified_true_benign)/ (total_trainingtest_sample)

Problem 2b

- Using the same setting(s) as developed in Problem 2a, run the classifier over all the entries in the blind sample (breast-cancer-wisconsin_mod_real.txt)
 - Produce a text file which contains only the ID of the samples which your classifier classifies as malignant (last_name.malignant_ID.txt)
 - Produce a text file which contains only the ID of the samples which your classifier classifies as benign (last_name.benign_ID.txt)
 - Basic text files. No Microsoft Word documents, Adobe PDF, or any other extraneous text editor formats. Only a single ID number per line in the text file that can be easily read by numpy.loadtxt().
 - Example online at http://www.nbi.dk/~koskinen/Teaching/
 AdvancedMethodsInAppliedStatistics2016/data/koskinen.benign_ID.txt
 - Any and all duplicates, i.e. two samples with the same ID, should be kept and included in the text files and analysis

• Data has the following probability distribution function G(x) with an unknown value of f:

$$G(x; f) = \frac{1}{N}(\cos(f \cdot x) + \frac{3}{x+1} + 1)$$
 for $1 \le x \le 10$ normalization is $N = \frac{\sin(10f) - \sin(f)}{f} + 9 + \ln(\frac{1331}{8})$

- The data is generated over a truncated range of $1 \le x \le 10$ and can be found at
 - (www.nbi.dk/~koskinen/Teaching/AdvancedMethodsInAppliedStatistics2016/data/ Exam_Prob3.txt)

Normalization from WolframlAlpha

Problem 3a

- We want to find f, but the likelihood as a function of f has many local minima/maxima. For this problem the search range will be constrained to $0 < f \le 20$ for simplicity.
- From other sources, there is a bayesian prior on the 'true' value of f. The prior is a normalized gaussian with a gaussian width of 0.5, i.e. $\sigma_f = 0.5$, centered at f=15.
- Using the above prior and the function from the previous slide, find the maximum a posteriori (MAP) value of f, i.e. mode of the bayesian posterior
 - Use a Markov Chain Monte Carlo technique
 - Remember: from Lec. 4 $\mathcal{L}(\theta|x) = P(x|\theta)$ and $\mathcal{L}(\theta) = \prod_{i=0}^{n} f(x_i;\theta)$

Problem 3b

- Using the same likelihood and prior from the previous slides for Problem 3, plot a histogram of the stable posterior distribution sampled points as a function of f
 - Whether to include or omit burn-in points/trials/steps is up to you
 - The histogram should include at least 500 sample points/trials/steps, but preferably less than 10001
 - Remember, the posterior distribution may have multiple local maxima/minima so you may end up with multiple bumps related to those minima/maxima

Problem 3c

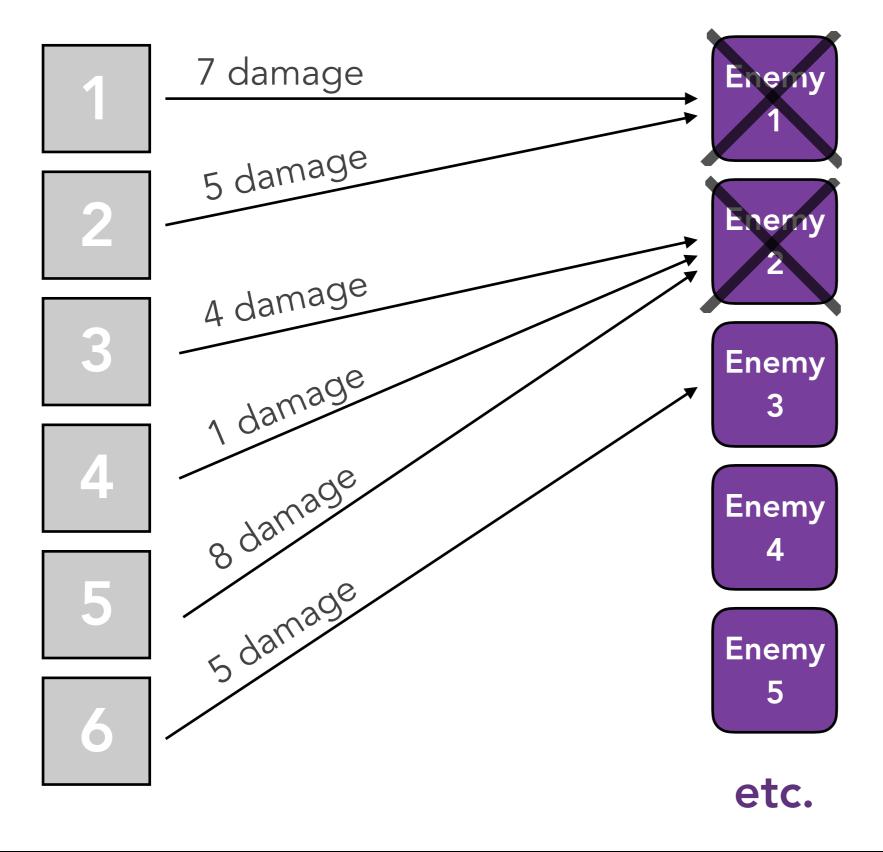
- Plot the posterior distribution using the same likelihood and the normalized gaussian prior center at f=15, but now the prior has a gaussian width of σ_f = 2.5
 - Does the posterior distribution change from when the gaussian width was 0.5?
 - If not describe why, and if so describe the changes.
 - What is the maximum a posterior value after including the new prior with a gaussian width of 2.5?

- You are playing a strategic turn-based computer game and you want to better understand likely outcomes. You have 6 units which are fighting 6+ enemies. In a turn, your units act only once and inflict damage in successive iteration to the first enemy in the queue, until the enemy has 0 or negative health, whereby that enemy is defeated. Once an enemy has been defeated, any of your remaining units which have not acted now inflict damage to the next enemy in the queue, and on and on until all your units have acted
 - Your units only individually act once during a turn to inflict damage
 - Damage inflicted follows a poisson likelihood

Problem 4a

- Find the mean number of enemies defeated in a single turn:
 - When the expected damage inflicted individually by each of your units is 5
 - Enemies can individually take 12 damage before being defeated
 - Your 6 units always individually inflict damage, i.e. any random samples of 0 should be rounded up to 1
 - An example illustration is on the next slide
- Make a histogram of the number of 'defeated enemies per turn' for 1000 unique and independent trials/turns
 - Each trial is a fresh set of enemies, i.e. for each trial all of the enemies should start w/ 12 health

Single Turn Example



In this example, individual enemies can receive 12 damage before being defeated

In total 2 enemies were defeated for this turn

Problem 4b

- Using the same values from 4a now include that your 6 units vary in individual accuracy and have some probability to inflict damage, or miss thereby inflicting no damage
 - The probability per unit to inflict damage is [90%, 80%, 60%, 90%, 60%, 70%]
 - Follow the order in the above array for calculations/plots
- Out of 5000 trials, what percentage of the time will no enemies be defeated in a turn, and what is the uncertainty on that percentage?

Problem 4c

- Using the same setup and values from 4b, test the new reorderings below, of your units inflicting damage versus the ordering in 4b of [90%, 80%, 60%, 90%, 60%, 70%]
 - Sorted ascending [60%, 60%, 70%, 80%, 90%, 90%]
 - Sorted descending [90%, 90%, 80%, 70%, 60%, 60%]
- Are the ascending and descending statistically compatible with the original ordering for the number of enemies defeated per turn?
 - Show and/or briefly explain your results

Select only 1 of the following problems for submission. Do all the parts.

- Apply the likelihood ratio test to the experiment from question 1 of Problem Set 2, with the PDF given below.
 - I have changed the previously poor notation to avoid some confusion, and 'T' is now replaced by 'b'

PDF =
$$f(t; b, \sigma_t) = \frac{1}{2b} \exp\left(\frac{\sigma_t^2}{2b^2} - \frac{t}{b}\right) \operatorname{erfc}\left(\frac{\sigma_t}{\sqrt{2}b} - \frac{t}{\sqrt{2}\sigma_t}\right)$$

• Neither b nor σ_t are explicitly known, and we want to test whether b=1 second can be rejected. We can do so via a hypothesis test, where the two hypotheses H₀ and H₁ are given as:

$$b_0 = 1.0 \ s$$

$$H_0: b = b_0$$

$$H_1: b \neq b_0$$

Problem 5 (cont.)

Use the likelihood ratio test:

$$\lambda = \frac{\mathcal{L}(\hat{\omega})}{\mathcal{L}(\hat{\Omega})} \qquad \qquad \omega \text{ given by } b = b_0, \ 0 < \sigma_t < \infty$$
$$\Omega \text{ given by } 0 < b < \infty, \ 0 < \sigma_t < \infty$$

- Where $\mathcal{L}(\hat{\omega})$ is the value of the null hypothesis likelihood calculated using the maximum likelihood estimator(s) $\hat{\omega}$
- Compute:

$$-2\ln\lambda = -2[\ln(L(\hat{\omega})) - \ln(L(\hat{\Omega}))]$$

if $\lambda \approx 1$, then the null hypothesis cannot be excluded if $\lambda \approx 0$, then the null hypothesis is unlikely true

Problem 5a

- There are 20000 events in the online file below, which corresponds to 100 simulated pseudo-experiments where each pseudo-experiment has 200 events.
- For each of the 100 pseudo-experiments find the values of the In-likelihoods that are maximized for the two hypotheses, i.e. $\ln(L(\hat{\omega}))$ and $\ln(L(\hat{\Omega}))$ and calculate -2ln(λ)
- As a histogram, plot the values of $-2\ln(\lambda)$
- The data is at http://www.nbi.dk/~koskinen/Teaching/
 AdvancedMethodsInAppliedStatistics2016/data/
 Exam_Prob5_NucData.txt

Problem 5b

- Assuming that $-2\ln(\lambda)$ is chi-squared distributed, how many pseudo-experiments of the 100 are expected to have $-2\ln(\lambda) > 2.706$?
- How many pseudo-experiments actually have $-2\ln(\lambda) > 2.706$?
- Bonus question: Why did I choose 2.706?

Problem 5c

• Using all 20000 events as a single pseudo-experiment, can the null hypothesis (H_0) be rejected at 3σ confidence?

- The artist Jackson Pollock was famous for creating paintings that look fractal, or scale invariant. On the next slide is an image of his piece "One: Number 31."
- A script has been used to convert the entire image to black and white and then write a single horizontal row of pixels into a text file at (http://www.nbi.dk/~koskinen/Teaching/AdvancedMethodsInAppliedStatistics2016/data/
 One_row.txt
 - Values are in 8-bit grayscale
 - 0 is black
 - 255 is white

Problem 6 - Photo



 Lower resolution image than the image posted on the class webpage (http://www.nbi.dk/~koskinen/Teaching/
 AdvancedMethodsInAppliedStatistics2016/data/One_Number31.png)

Problem 6 (cont.)

- Wavelet coefficients are labelled by two indices: a scale (or level), and a positional index. The RMS of those coefficients that all have the same scale index provides a measure of the activity in a signal at that given scale.
- Enlarging a small section of a scale-invariant signal will result in a new signal that has the same characteristics as the original signal from which the small piece was taken. This means that the size of fluctuations in a scale-invariant signal must be proportional to the scale of the signal.

Problem 6a

- Take the single row of Jackson Pollock data in the file below and plot the RMS of the wavelet coefficients vs. the scale of the wavelet coefficients.
 - (http://www.nbi.dk/~koskinen/Teaching/ AdvancedMethodsInAppliedStatistics2016/data/One_row.txt)
- What would you expect this plot to look like if the signal were purely scale-invariant?
- Does Pollock's artwork deviate from scale invariance, and if so, how?

Problem 6b

- White noise is not scale invariant because the size of fluctuations is constant vs. scale, meaning the power spectrum is flat in the frequency domain. However, other types of noise are scale invariant, an example being the sound of the sea.
- Use the Haar and D4 wavelet bases to generate two scale-invariant noise samples, each 256 bins long. What are the differences between these two samples?
- Bonus Question: A frequent cigarette smoker while working, many Pollock paintings have cigarette stubs embedded in the paint. Would it be worthwhile to develop a "Pollock" wavelet bases to account for non-paint debris?