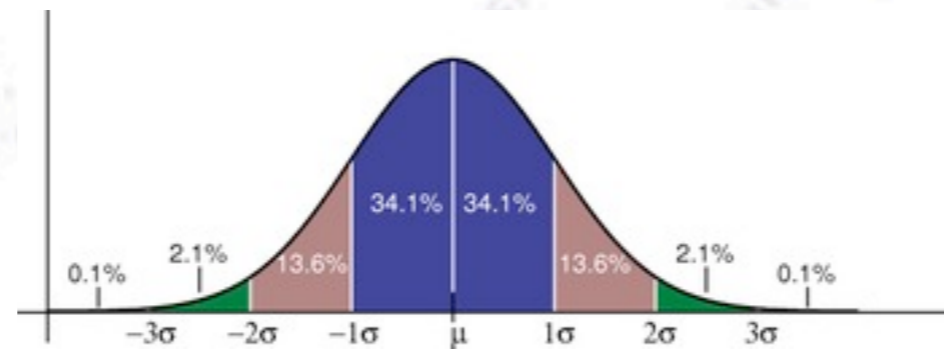


# Applied Statistics

## Principle of maximum likelihood



Troels C. Petersen (NBI)



*"Statistics is merely a quantization of common sense"*

# Likelihood function



*“I shall stick to the principle of likelihood...”*  
[Plato, in Timaeus]



# Likelihood function



Given a PDF  $f$ , what is the chance that with  $N$  observations,  $x_i$  falls in the interval  $[x_i, x_i + dx_i]$ ?

$$\mathcal{L}(\theta) = \prod_i f(x_i, \theta) dx_i$$



# Likelihood function

Given a set of measurements  $\mathbf{x}$ , and parameter(s)  $\theta$ , the likelihood function is defined as:

$$\mathcal{L}(x_1, x_2, \dots, x_N; \theta) = \prod_i p(x_i, \theta)$$

The **principle of maximum likelihood** for parameter estimation consist of maximizing the likelihood of parameter(s) (here  $\theta$ ) given some data (here  $\mathbf{x}$ ).

The likelihood function plays a central role in statistics, as it can be shown to be:

- Consistent (converges to the right value!)
- Asymptotically normal (converges with Gaussian errors).
- Efficient (reaches the Cramer-Rao lower bound for large N).

To some extend, this means that the likelihood function is “optimal”, that is, if it can be applied in practice.





# Likelihood vs. Chi-Square

For computational reasons, it is often much easier to minimize the logarithm of the likelihood function:

$$\frac{\partial \ln \mathcal{L}}{\partial \theta} \bigg|_{\theta = \bar{\theta}} = 0$$

In problems with Gaussian errors, it turns out that the **likelihood function** boils down to the **Chi-Square** with a constant offset and a factor -2 in difference.

In practice, the likelihood comes in two versions:

- Binned likelihood (using Poisson).
- Unbinned likelihood (using PDF).

The “trouble” with the likelihood is, that it is unlike the Chi-Square, there is NO simple way to obtain a probability of obtaining certain likelihood value!



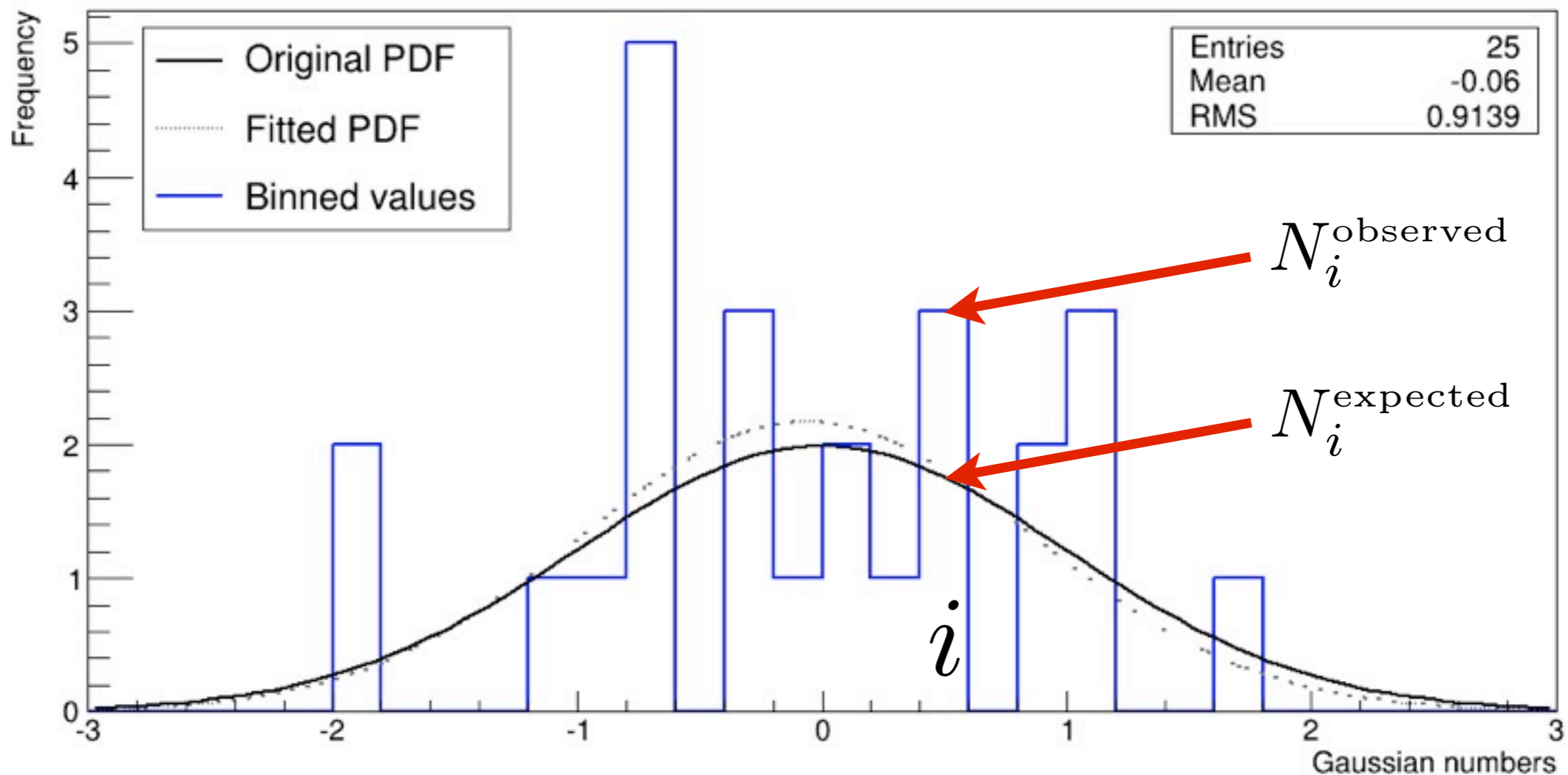
# Binned Likelihood

The binned likelihood is a sum over bins in a histogram:

$$\mathcal{L}(\theta)_{\text{binned}} = \prod_i^{N_{\text{bins}}} \text{Poisson}(N_i^{\text{expected}}, N_i^{\text{observed}})$$

$$f(n; \nu) = \frac{\nu^n}{n!} e^{-\nu}$$

Distribution of 25 unit Gaussian numbers





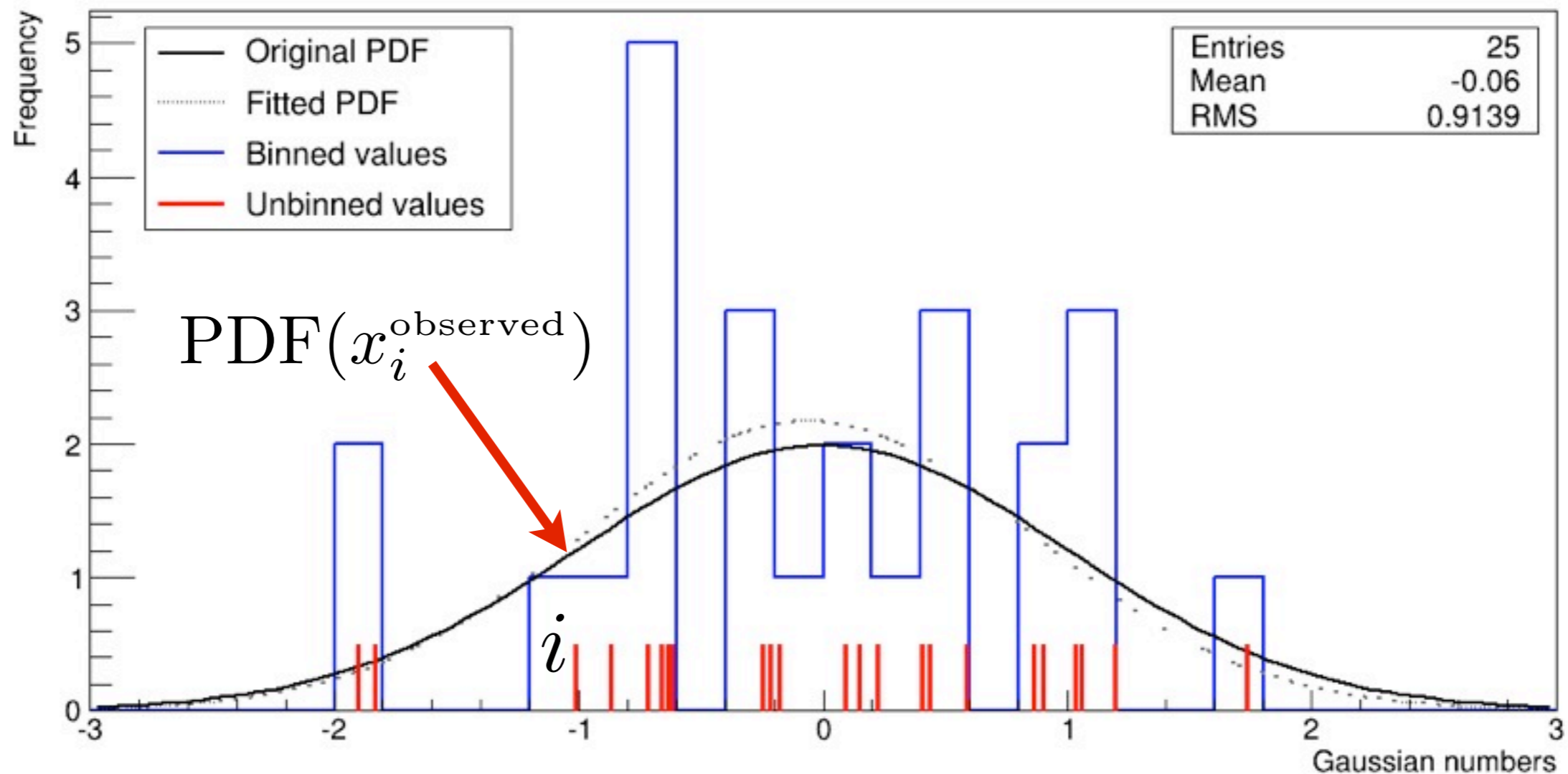


# Unbinned Likelihood

The binned likelihood is a sum over single measurements:

$$\mathcal{L}(\theta)_{\text{unbinned}} = \prod_i^{N_{\text{meas.}}} \text{PDF}(x_i^{\text{observed}})$$

Distribution of 25 unit Gaussian numbers



# Notes on the likelihood

For a large sample, the maximum likelihood (ML) is indeed unbiased and has the minimum variance - that is hard to beat! However...

For the ML, you have to know your PDF. This is also true for the Chi-Square, but unlike for the Chi-Square, you get no goodness-of-fit measure!

Also, the small statistics, the ML is not necessarily unbiased (nor is the ChiSquare)! Careful with this. The way to avoid this problem is using simulation - more to follow.





# The likelihood ratio test

Not unlike the Chi-Square, where one can compare  $\chi^2$  values, the likelihood between two competing hypothesis can be compared (SAME offset constant/factor!).

While their individual LLH values do not say much, their RATIO says everything!

As with the likelihood, one often takes the logarithm and multiplies by -2 to match the Chi-Square, thus the “test statistic” becomes:

$$\begin{aligned} D &= -2 \ln \left( \frac{\text{likelihood for null model}}{\text{likelihood for alternative model}} \right) \\ &= -2 \ln(\text{likelihood for null model}) + 2 \ln(\text{likelihood for alternative model}) \end{aligned}$$

If the two hypothesis are simple (i.e. no free parameters) then the **Neyman-Pearson Lemma** states that this is the best possible test one can make.

If the alternative model is not simple, this difference behaves like a Chi-Square distribution with  $N_{\text{dof}} = N_{\text{dof}}(\text{alternative}) - N_{\text{dof}}(\text{null})$